

# An enhanced variance estimator with power transformation

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## ABSTRACT

In this study, we propose an enhanced variance estimator with power transformation to estimate the population variance of the study variable. The properties (Bias and Mean Square Error) of the proposed estimator were derived up to the first order of approximation using the Taylor series technique. The conditions for the proposed estimator to be better than existing related estimators are established. The empirical study was conducted using two natural populations and the result revealed that the proposed estimator is more efficient.

**Keywords:** Population variance, Sampling, Auxiliary variable, Mean Square Error, Bias

## 1. INTRODUCTION

The ratio method of estimation for the finite population was developed by (Isaki, 1983) in the presence of population variance of the auxiliary variable. This estimation procedure is used when there is a positive correlation between the study and the auxiliary variable. Auxiliary information is very useful for the enhancement of an estimator and that is why authors used it often during the development of estimators, in this case (Audu et al., 2016; Singh et al., 2007; Singh et al., 2011; Ahmed et al., 2003; Gupta and Shabbir, 2008; Singh and Solanki, 2013; Solanki and Singh, 2013; Adejumobi et al., 2022; Upadhyaya and Singh, 1999; Yadav and Kadilar, 2013; Kadilar and Cingi, 2006; Subramani and Kumarapandiyan, 2012; Adejumobi and Yunusa, 2022).

Consider  $\varphi = (1, 2, \dots, N)$  as a population of size  $N$  and  $X, Y$  be auxiliary and study variable having values  $(X_i, Y_i) \in R > 0$  on the  $i^{\text{th}}$  unit of  $\varphi_i (1 < i \leq N)$ , we assume a positive correlation  $\rho > 0$  between the auxiliary and study variables. Let  $\sigma_x^2$  be the finite population variance of  $X$  and  $\sigma_y^2$  be the sample variance based on the random sample of size  $n$  chosen without replacement.

Let us define the following relation for error terms as

$$e_0 = \frac{s_y^2 - S_y^2}{S_y^2}, \quad e_1 = \frac{s_x^2 - S_x^2}{S_x^2} \quad \text{such that}$$

$$s_y^2 = S_y^2 (1 + e_0), \quad s_x^2 = S_x^2 (1 + e_1)$$

$$E(e_0) = E(e_1) = 0, \quad E(e_0^2) = \gamma(\lambda_{40} - 1), \quad E(e_1^2) = \gamma(\lambda_{04} - 1),$$

$$E(e_0 e_1) = \gamma(\lambda_{22} - 1), \quad \gamma = \frac{1}{n}$$

In this study, we proposed an enhanced variance estimator with power transformation for the estimation of population variance of the study variable.

### Estimators in literature

The usual unbiased variance estimator is given by:

$$t_0 = s_y^2 = (n-1)^{-1} \sum_{i=1}^n (y_i - \bar{y})^2 \quad (1)$$

The variance is given as:

$$\text{var}(t_0) = \gamma S_y^4 (\lambda_{40} - 1) \quad (2)$$

Isaki, (1983) proposed the ratio estimator for population variance as

$$t_1 = s_y^2 \left( \frac{S_x^2}{S_x^2} \right) \quad (3)$$

The mean square error is given by

$$MSE(t_1) = \gamma S_y^4 [(\lambda_{40} - 1) + (\lambda_{04} - 1) - 2(\lambda_{22} - 1)] \quad (4)$$

Singh et al., (2009) proposed a ratio-type exponential estimator for the estimation of finite population variance as

$$MSE(t_2) = \gamma S_y^4 \left[ (\lambda_{40} - 1) + \frac{(\lambda_{04} - 1)}{4} - (\lambda_{22} - 1) \right] \quad (5)$$

## 2. MATERIALS AND METHODS

### Proposed estimator

Having studied the estimators in literature and following estimation strategy of (Yunusa et al., 2021) population mean estimator, we proposed an enhanced variance estimator in the form

$$t_{AM} = 2^{-1} s_y^2 \left[ \left( \frac{S_x^2}{S_x^2} \right)^\alpha + \exp \left( \frac{S_x^2 - s_x^2}{S_x^2 + s_x^2} \right) \right] \quad (6)$$

where,  $\alpha$  is the constant to be determined.

Express  $t_{AM}$  in terms of error  $e_i^s$  ( $i = 0, 1$ ) we have

$$t_{AM} = 2^{-1} S_y^2 (1 + e_0) \left[ \left( \frac{S_x^2 (1 + e_1)}{S_x^2} \right)^\alpha + \exp \left( \frac{S_x^2 - S_x^2 (1 + e_1)}{S_x^2 + S_x^2 (1 + e_1)} \right) \right] \quad (7)$$

Simplify (7) we have

$$t_{AM} = 2^{-1} S_y^2 (1 + e_0) \left[ (1 + e_1)^\alpha + 1 - \frac{e_1}{2} + \frac{3e_1^2}{8} \right] \quad (8)$$

$$t_{AM} = \frac{S_y^2 (1 + e_0)}{2} \left[ 1 + \alpha e_1 + \frac{\alpha(\alpha-1)e_1^2}{2} + 1 - \frac{e_1}{2} + \frac{3e_1^2}{8} \right] \quad (9)$$

$$t_{AM} = S_y^2 (1 + e_0) \left[ 1 + \left( \alpha - \frac{1}{2} \right) e_1 + \left( \frac{\alpha^2 - \alpha}{2} + \frac{3}{8} \right) e_1^2 \right] \quad (10)$$

Subtract  $S_y^2$  from both sides and simplify to first order of approximation, we have

$$t_{AM} - S_y^2 = S_y^2 \left[ e_0 + \left( \alpha - \frac{1}{2} \right) e_1 + \left( \alpha - \frac{1}{2} \right) e_0 e_1 + \left( \frac{\alpha^2 - \alpha}{2} + \frac{3}{8} \right) e_1^2 \right] \quad (11)$$

$$Bias(t_{AM}) = S_y^2 \gamma \left[ \left( \alpha - \frac{1}{2} \right) (\lambda_{22} - 1) + \left( \frac{\alpha^2 - \alpha}{2} + \frac{3}{8} \right) (\lambda_{04} - 1) \right] \quad (12)$$

Squaring and taking expectation of equation (11), we obtain the mean square error up to first order of approximation as:

$$MSE(t_{AM}) = \gamma S_y^4 \left[ (\lambda_{40} - 1) + \left( \alpha - \frac{1}{2} \right)^2 (\lambda_{04} - 1) + 2 \left( \alpha - \frac{1}{2} \right) (\lambda_{22} - 1) \right] \quad (13)$$

Differentiating (13) with respect to  $\alpha$ , equate to zero and solve for  $\alpha$ , we obtain

$$\alpha = \frac{1}{2} - \frac{(\lambda_{22} - 1)}{(\lambda_{04} - 1)} \quad (14)$$

Substituting (14) into (13) to obtain the minimum mean square error (MSE) of  $t_{AM}$  is

$$MSE(t_{AM})_{\min} = \gamma S_y^4 \left[ (\lambda_{40} - 1) - \frac{(\lambda_{22} - 1)^2}{(\lambda_{04} - 1)} \right] \quad (15)$$

### Efficiency Comparison

In this section, the estimator  $t_{AM}$  is more efficient than  $t_0, t_1$  and  $t_2$ , if the following conditions are satisfied

$$MSE(t_{AM})_{\min} < \text{var}(t_0) \text{ if}$$

$$\frac{(\lambda_{22} - 1)^2}{(\lambda_{04} - 1)} > 0 \quad (16)$$

$$MSE(t_{AM})_{\min} < MSE(t_1) \text{ if}$$

$$(\lambda_{04} - 1) > (\lambda_{22} - 1) \left( 2 - \frac{(\lambda_{22} - 1)}{(\lambda_{04} - 1)} \right) \quad (17)$$

$$MSE(t_{AM})_{\min} < MSE(t_2) \text{ if}$$

$$(\lambda_{04} - 1) > (\lambda_{22} - 1) \left( 1 - \frac{(\lambda_{22} - 1)}{(\lambda_{04} - 1)} \right) \quad (18)$$

### 3. EMPIRICAL STUDY

In this section, in order to elucidate the performance of the proposed estimator efficiency over  $t_0, t_1$  and  $t_2$ , we use two natural populations.

Table 2 shows the mean square errors (MSEs) and percentage relative efficiencies (PREs) of the proposed estimator and existing related estimators and. It can be observed that the proposed estimator has the minimum mean square error (min MSE) and higher percentage relative efficiency. This reveals that it is more efficient than its counterparts.

**Table 1** Dataset used for the empirical study

Parameters	Population 1	Population 2
$N$	34	69
$n$	15	40
$\bar{X}$	208.88	4591.07
$\bar{Y}$	199.44	4514.89
$C_x$	0.72	1.38
$C_y$	0.75	1.35
$\lambda_{40}$	3.6161	7.66
$\lambda_{04}$	2.8266	9.84
$\lambda_{22}$	3.0133	8.19
$\rho_{yx}$	0.98	0.96

**Table 2** Mean Square Error of the proposed and existing estimators

Estimators	Population 1		Population 2	
	MSE	PRE	MSE	PRE
$t_0$	873086634.77	100.00	2.29794E+14	100.00
$t_1$	13886748.57	628.72	3.8644E+13	594.64
$t_2$	19758683.55	441.87	2.4363E+13	396.43
$t_3^d$	13249882.6	658.94	2.80177E+13	820.17

## 4. CONCLUSION

In this study, we proposed an enhanced variance estimator and derived its properties (Bias and MSE) as well as efficiency conditions of  $t_{AM}$  being efficient than  $t_0, t_1$  and  $t_2$ . The empirical study revealed that  $t_{AM}$  is more efficient with the evident of minimum MSE and higher PRE. Hence, the proposed estimator  $t_{AM}$  is recommended in practical use.

### Informed consent

Not applicable.

### Ethical approval

Not applicable.

### Conflicts of interests

The authors declare that there are no conflicts of interests.

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### Data and materials availability

All data associated with this study are present in the paper.

## REFERENCES AND NOTES

1. Adejumobi A, Yunusa MA, Audu A. Improved Modified Classes of Regression Type Estimators of Finite Population Mean in the Presence of Auxiliary Attribute. *Orient J Phys Sci* 2022; 07:41-47. doi: 10.13005/OJPS07.01.07
2. Adejumobi A, Yunusa MA. Some Improved Class of Ratio Estimators for Finite Population Variance with the Use of Known Parameters. *LC Int J STEM* 2022; 3(3):2708-7123. doi: 10.5281/zenodo.7271392
3. Ahmed MS, Dayyeh WA, Hurrairah AAO. Some estimators for finite population variance under two-phase sampling. *Stat Transit* 2003; 6(1):143-150.
4. Audu A, Adewara AA, Singh RVK. Improved class of ratio estimators for finite population variance. *Glob J Sci Front Res Mathema Decis Sci* 2016; 16(2):17-24.
5. Gupta S, Shabbir J. Variance estimation in simple random sampling using auxiliary variable information. *Hacet J Math Stat* 2008; 37:57-67.
6. Isaki CT. Variance estimation using auxiliary information. *J Am Stat Assoc* 1983; 78:117-123.
7. Kadilar C, Cingi H. Ratio estimators for population variance in simple and stratified sampling. *Appl Math Comput* 2006; 173:1047-1059.
8. Khoshnevisan M, Singh R, Chauhan P, Sawan N, Smarandache F. A general family of estimators for estimating population variance using known value of some population parameters. *Far East J Theor Stat* 2007.
9. Singh HP, Solanki RS. A new procedure for variance estimation in simple random sampling using auxiliary information. *Stat Pap* 2013; 54(2):479-497.
10. Singh R, Chauhan P, Sawan N, Smarandache F. Improved exponential estimator for population variance using two auxiliary variables. *Ital J Pure Appl Math* 2011; 28:101-108.
11. Singh R, Chauhan P, Sawan N, Smarandache F. Improved exponential estimator for population variance using two auxiliary variables. *Octogon Math Mag* 2009; 17(2):647-674.
12. Solanki RS, Singh HP. An improved class of estimators for the population variance. *Model Assist Stat Appl* 2013; 8(3):229-238.
13. Subramani J, Kumaranpandiyan G. Variance estimation using quartiles and their functions of an auxiliary variable. *Int J Stat Appl* 2012; 2(5):67-72.
14. Upadhyaya LN, Singh HP. An estimator for population variance that utilizes the kurtosis of an auxiliary variable in sample survey, *Vikram Math J* 1999; 19:14-17.
15. Yadav SK, Kadilar C. Improved exponential type ratio estimator of population variance. *Rev Colomb Estad* 2013; 36 (1):45-152.
16. Yunusa MA, Audu A, Musa N, Beki DO, Rashida A, Bello AB, Hairullahi MU. Logarithmic ratio-type estimator of population coefficient of variation. *Asian J Probab Stat* 2021; 14(2):13-22.